

Prognostic significance of NPM1 mutations in acute myeloid leukemia: A meta-analysis

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Received August 2, 2013; Accepted November 4, 2013

DOI: 10.3892/mco.2013.222

Abstract. Nucleophosmin 1 (NPM1) mutations have been identified in a substantial number of patients with acute myeloid leukemia (AML). Favorable outcomes in AML cases with NPM1 mutations have been previously reported. However, widely differing survival estimates have been indicated. Therefore, a meta-analysis of nine studies including a total of 4509 subjects was performed. The frequency of NPM1 mutations was found to be 6.45-56.08%. NPM1-mutation type (NPM1-mt) patients had >2-fold higher odds of achieving complete remission compared with NPM1-wild-type (NPM1-wt). The summary hazard ratio (HR) of NPM1-mt/NPM1-wt for disease-free survival (DFS) and OS was 0.67 and 0.63, respectively. In conclusion, these findings suggest that the NPM1 mutation has a favorable effect on the outcome for AML. The present meta-analysis was based on data abstracted from observational studies. However, the results obtained may justify the risk-adapted therapeutic strategies for AML according to the NPM1 status.

Introduction

Acute myeloid leukemia (AML) is a heterogeneous clonal disorder characterized by autonomous proliferation and impaired differentiation of hematopoietic progenitor cells (1,2). It is the most common malignant myeloid disorder in adults. Cytogenetic aberrations and molecular genetic alterations provide significant prognostic information for determining the response to chemotherapy and survival outcome (3,4). An increasing number of genetic abnormalities revealed in AML

have also contributed to our understanding of the mechanisms and process of leukemogenesis, which leads to an improvement of risk-stratification, and the development of individualized therapies and response assessment (5).

Among these genetic alterations, a potential prognostic genetic marker is the nucleophosmin 1 (NPM1) gene, which is important in many tumor-associated chromosomal translocations (6). NPM1 is an ubiquitously expressed phosphoprotein and continuously shuttles between the cytoplasm and nucleus (7,8). Several functions for this protein have been described, including the binding of p53 (9), the initiation of centrosome duplication (10), and ribosomal protein assembly and transport (11). More recently, NPM1 exon 12 mutations have been reported to be involved in leukemogenesis, and detected in ~35% of AML cases (12,13). However, the prognostic implications of NPM1 mutations are less clear and are notably variable among different institutions (13-18). A study conducted by Konoplev *et al* including 252 AML patients suggested NPM1 mutations did not impact overall survival (OS) and event-free survival (EFS) (17). However, findings of previous studies have indicated that the NPM1 mutation has a favorable effect on the outcome for AML (13-16,19-22). For this reason, we performed an updated meta-analysis of 9 published studies in order to investigate the prognostic significance of NPM1 mutations for AML.

Materials and methods

Selection of studies. Studies were eligible for inclusion in the meta-analysis if they were: i) original articles written in English and published up to January 2013; ii) dealt only with untreated AML patients; iii) offered survival information based on the NPM1 status, including, NPM1 mutations and NPM1-wild-type (NPM1-wt), and iv) provided survival information on response to induction therapy, including, complete remission (CR), disease-free survival (DFS) and/or OS. Studies were excluded if they focused exclusively on acute promyelocytic leukemia. Multiple reports of a single study were considered as one publication, and only the most recent article was examined.

A computerized literature search of the PubMed, Medline and EMBASE databases was conducted using the free text search term AML AND nucleophosmin AND survival, with the publication period limited to prior to January 2013, and the language to English. The initial search yielded a total of

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Key words: acute myeloid leukemia, nucleophosmin 1, mutation, prognosis, meta-analysis

Table I. List of studies included in the meta-analysis.

| Author (Refs.) | Publication year | Region | Subjects (n) | NPM1 mutation (%) | Normal karyotype (%) |
|------------------------------|------------------|-------------------|--------------|-------------------|----------------------|
| Thiede <i>et al</i> (14) | 2006 | Germany | 1485 | 27.47 | 79.41 |
| Döhner <i>et al</i> (15) | 2005 | Germany and USA | 300 | 48.33 | NR |
| Schnittger <i>et al</i> (16) | 2005 | Germany and Italy | 401 | 52.90 | 100.00 |
| Suzuki <i>et al</i> (19) | 2005 | Japan | 190 | 30.53 | 63.79 |
| Verhaak <i>et al</i> (13) | 2005 | The Netherlands | 275 | 34.55 | 77.89 |
| Mullighan <i>et al</i> (18) | 2007 | USA | 93 | 6.45 | 50.00 |
| Gale <i>et al</i> (20) | 2008 | United Kingdom | 1217 | 41.33 | 67.59 |
| Becker <i>et al</i> (21) | 2010 | Germany and USA | 148 | 56.08 | NR |
| Boonthimat <i>et al</i> (22) | 2008 | Thailand | 400 | 26.25 | 86.67 |

NPM1, nucleophosmin 1; NR, not reported.

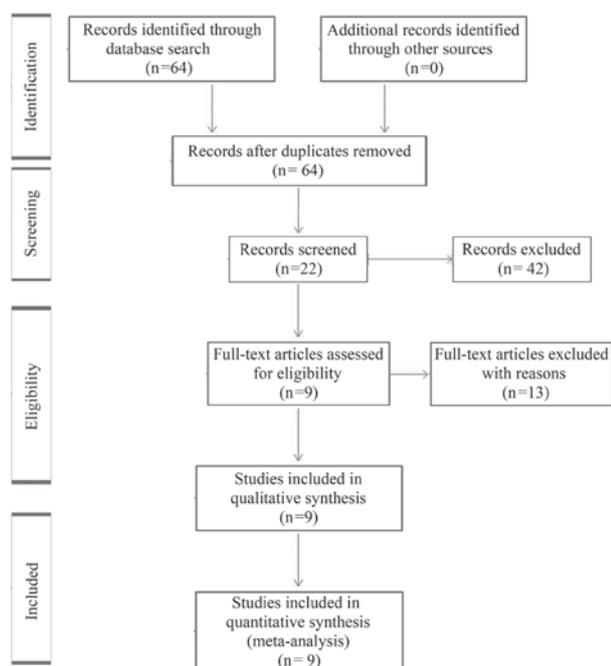


Figure 1. Flow diagram showing the process of identifying and selecting relevant studies.

64 articles, and the titles and abstracts of these papers were reviewed, resulting in the exclusion of 42 articles, with 22 candidate articles. Of the 22 articles, 13 full-text articles were excluded as survival information was unavailable. In total, 9 studies satisfied the eligibility criteria and were included in the meta-analysis (Table I). The reasons for excluding articles are shown in Fig. 1 (23-35).

Data extraction and quality assessment. In order to avoid bias in the data abstraction process, the reviewers Y.F. Liu and P.C. He independently retrieved the data from the articles and subsequently compared the results. All data were assessed for internal consistency and disagreements were resolved by discussion. Characteristics abstracted from the articles included the name of the first author, year of publication, location of

the study, number of subjects, mean or median values of age and median white blood cell (WBC) counts, the incidence of NPM1 mutations, percentage of cases with normal karyotype, outcomes including hematologic CR rate, hazard ratio (HR) and 95% confidence interval (CI) for DFS and OS according to the NPM1 status based on multivariate analysis. When the data required for the analysis could not be abstracted, attempts were made to contact the investigators who conducted the studies.

The quality of evidence and the strength of recommendations were evaluated by GRADE profiler (version 3.2) (36). Any discrepancies in quality assessments were resolved by consensus among the authors. The overall quality of the data was graded as moderate.

Quantitative data synthesis. HR was used to assess the survival effect of NPM1 mutations compared with wild-type. The natural logarithm of a crude HR and its variance within the study was calculated by using the abstracted survival probabilities at each time point with the methods proposed by Parmar *et al* (37), and those described previously (38). HR was calculated to show how many times higher the probability of survival failure was for the patients with NPM1 mutations compared with those with wild-type, as an HR less than unity suggests that NPM1 mutations yield a better survival rate compared with wild-type.

The odds ratio (OR) was calculated to describe the probability of CR following induction therapy based on NPM1 mutation status. An OR >1 indicates that patients with NPM1 mutations are associated with an improved CR rate compared with those without the mutations.

A DerSimonian Laird random method was used to calculate summary HR or OR and their 95% CI. Initially, the fixed effect and random-effect models were used to calculate summary HRs, but eventually the random-effect model was selected. Begg's funnel plots (39) and Egger's test (40) were used to detect possible publication bias. The between-study variation (τ^2) from the Q statistic was also calculated (41).

Statistical analysis. Statistical analyses were performed using STATA 12 software (College Station, TX, USA). $P < 0.05$ was considered to indicate a statistically significant test result for summary HR or OR.

Table II. Diagnostic characteristics according to the NPM1 status in the AML patients.

| Author (Refs.) | NPM1 status | Subjects (n) | Age (years) | Median WBC count (10 ⁹ /l) | FLT3-ITD mutation (%) |
|------------------------------|-------------|--------------|------------------|---------------------------------------|-----------------------|
| Thiede <i>et al</i> (14) | NPM1-wt | 1077 | 58.0 (15.0-87.0) | 0.3-465 | 13.74 ^a |
| | NPM1-mt | 408 | 60.0 (18.0-83.0) | 0.5-380 | 40.20 |
| Döhner <i>et al</i> (5) | NPM1-wt | 155 | 16.0-60.0 | 0.41-369 | 24.52 ^a |
| | NPM1-mt | 145 | 18.0-60.0 | 0.2-345 | 40.69 |
| Schnittger <i>et al</i> (16) | NPM1-wt | 189 | 58.1 | 10.0 (0.1-361) | NR |
| | NPM1-mt | 212 | 55.8 | 38.7 (0.2-486) | NR |
| Suzuki <i>et al</i> (19) | NPM1-wt | 141 | 47.0 (15.0-85.0) | 23.3 (0.9-337.6) | 11.35 ^a |
| | NPM1-mt | 49 | 58.0 (15.0-77.0) | 52.2 (1.0-372) | 55.10 |
| Verhaak <i>et al</i> (13) | NPM1-wt | 180 | 39.7±13.3 | NA | 16.76 ^a |
| | NPM1-mt | 95 | 47.3±10.7 | NA | 49.47 |
| Mullighan <i>et al</i> (18) | NPM1-wt | 87 | 9.7 | 38.0 | NR |
| | NPM1-mt | 6 | 14.6 | 35.3 | 16.67 |
| Gale <i>et al</i> (20) | NPM1-wt | 714 | 41.0 | 18.5 | 19.19 ^a |
| | NPM1-mt | 503 | 46.0 | 35.4 | 41.35 |
| Becker <i>et al</i> (21) | NPM1-wt | 65 | 71.0 (60.0-83.0) | 7.0 (1.0-434.1) | 20.00 ^a |
| | NPM1-mt | 83 | 67.0 (60.0-81.0) | 26.2 (1.0-249.3) | 39.76 |
| Boonthimat <i>et al</i> (22) | NPM1-wt | 295 | 40.0 | 25.4 | 20.68 ^a |
| | NPM1-mt | 105 | 51.0 | 47.0 | 43.81 |

NPM1, nucleophosmin 1; AML, acute myeloid leukemia; WBC, white blood cell; NR, not reported; NA, not assessed; NPM1-mt, NPM1-mutation type; NPM1-wt, NPM1-wild-type. Statistically significant difference (^aP<0.05).

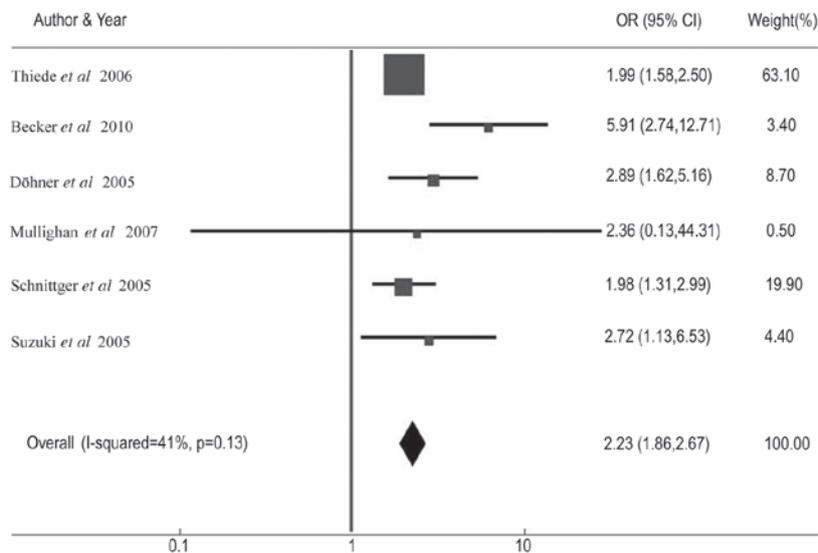


Figure 2. Forest plots of the odds ratio (OR) and 95% confidence intervals for complete remission. The size of the blocks or diamonds represents the weight for the random-effect model in the meta-analysis. OR>1 indicates that the presence of NPM1 mutations is associated with a higher complete remission rate.

Results

Study characteristics. As shown in Table I, 9 studies including a total of 4,509 subjects (2,903 with NPM1-wt, 1,606 with NPM1 mutations) were included in the meta-analysis. Four studies originated from Europe (13,14,16,20), two from Asia (19,22)

and three from the USA, two of which contained information pertaining to Germany (15,18,21). The frequency of NPM1 mutations varied between 6.45 and 56.08% for AML patients. NPM1 mutations were associated with a higher frequency of FLT3-ITD mutations in 7 studies (Table II) (13-15,19-22). The frequency of the normal karyotype was higher among NPM1

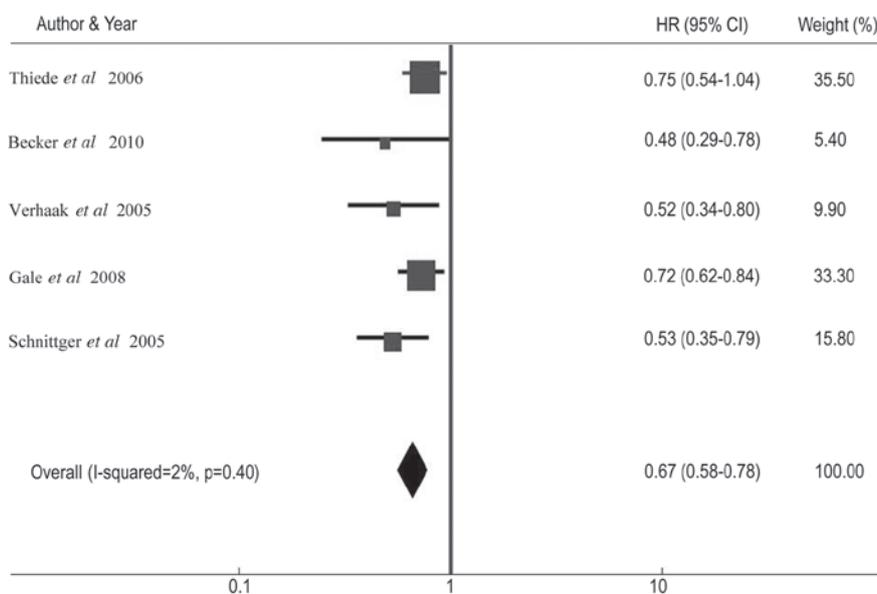


Figure 3. Forest plots of the hazard ratio (HR) and 95% confidence intervals (CIs) for disease-free survival. The size of the blocks or diamonds represents the weight for the random-effects model in the meta-analysis. HR<1 indicates that the presence of nucleophosmin 1 (NPM1) mutations is associated with an improved prognosis.

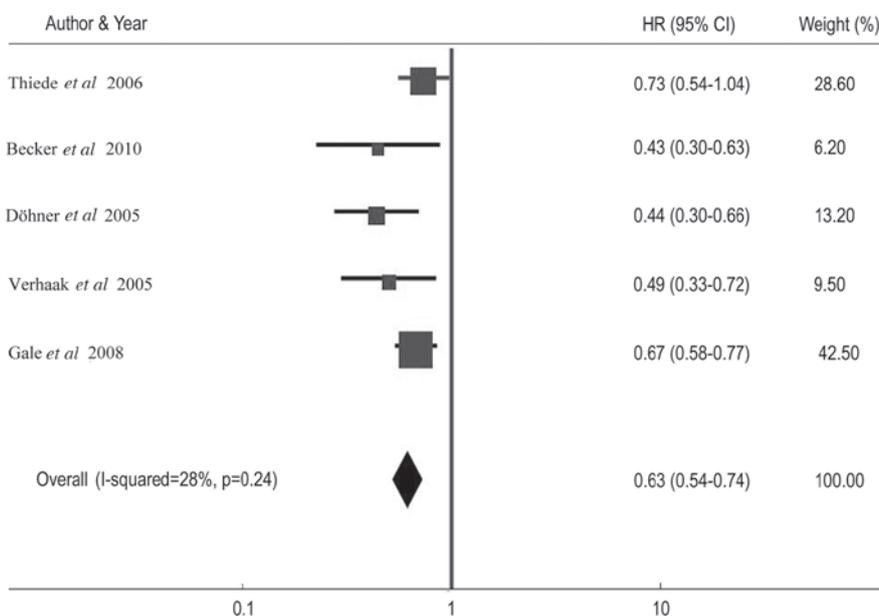


Figure 4. Forest plots of the hazard ratio (HR) and 95% confidence intervals (CIs) for overall survival. The size of the blocks or diamonds represents the weight for the random-effect model in the meta-analysis. HR<1 indicates that the presence of nucleophosmin 1 (NPM1) mutations is associated with an improved prognosis.

mutant patients (Table I) (13,14,16,18-20,22). No graphical or statistical evidence of publication bias for DFS or OS was identified.

Treatment outcomes. Table III shows the CR rate and HR for DFS and OS among AML patients with NPM1 mutations compared with patients without the mutations in individual studies. The summary OR for CR in the NPM1 mutant group was 2.23 (95% CI: 1.86-2.67, P<0.001) (Fig. 2). The summary HR for DFS of NPM1-mt/NPM1-wt was

0.67 (95% CI: 0.58-0.78; P<0.001) (Fig. 3), and the overall HR for OS of NPM1-mt/NPM1-wt was 0.63 (95% CI: 0.54-0.74; P<0.001) (Fig. 4). The test for heterogeneity, which evaluates the variation in study outcomes between studies in a meta-analysis, showed no significant heterogeneity among studies included in the DFS analysis (Q=4.08, df=4, P=0.40, $\tau^2=2$) and OS analysis (Q=5.52, df=4, P=0.24, $\tau^2=28$).

Furthermore, we performed a sensitivity test during the meta-analysis. Exclusion of any single study did not affect the overall results.

Table III. NPM1 mutations and outcomes in acute myeloid leukemia.

| Author (Refs.) | NPM1 status | Subjects (n) | CR (%) | HR for OS | 95% CI for OS | HR for DFS | 95% CI for DFS |
|------------------------------|-------------|--------------|--------------------|-----------|---------------|------------|----------------|
| Thiede <i>et al</i> (14) | NPM1-wt | 1077 | 41.60 ^a | 1.00 | Reference | 1.00 | Reference |
| | NPM1-nt | 408 | 58.60 | 0.73 | 0.58-0.91 | 0.75 | 0.54-1.04 |
| Döhner <i>et al</i> (15) | NPM1-wt | 155 | 68.50 ^a | 1.00 | Reference | NR | NR |
| | NPM1-nt | 145 | 86.00 | 0.44 | 0.30-0.66 | NR | NR |
| Schnittger <i>et al</i> (16) | NPM1-wt | 189 | 54.70 ^a | NR | NR | 1.00 | Reference |
| | NPM1-nt | 212 | 70.50 | NR | NR | 0.53 | 0.35-0.79 |
| Suzuki <i>et al</i> (19) | NPM1-wt | 141 | 68.80 ^a | NR | NR | NR | NR |
| | NPM1-nt | 49 | 85.70 | NR | NR | NR | NR |
| Verhaak <i>et al</i> (13) | NPM1-wt | 180 | NA | 1.00 | Reference | 1.00 | Reference |
| | NPM1-nt | 95 | NA | 0.49 | 0.33-0.72 | 0.52 | 0.34-0.80 |
| Mullighan <i>et al</i> (18) | NPM1-wt | 87 | 85.06 | NA | NA | NA | NA |
| | NPM1-nt | 6 | 100.00 | NA | NA | NA | NA |
| Gale <i>et al</i> (20) | NPM1-wt | 714 | NA | 1.00 | Reference | 1.00 | Reference |
| | NPM1-nt | 503 | NA | 0.67 | 0.58-0.77 | 0.72 | 0.62-0.84 |
| Becker <i>et al</i> (21) | NPM1-wt | 65 | 47.69 ^a | 1.00 | Reference | 1.00 | Reference |
| | NPM1-nt | 83 | 84.34 | 0.43 | 0.30-0.63 | 0.48 | 0.29-0.78 |

NPM1, nucleophosmin 1; CR, complete remission; HR, hazard ratio; 95% CI, 95% confidence interval; OS, overall survival; NR, not reported; NA, not assessed; DFS, disease-free survival; NPM1-nt, NPM1-mutation type; NPM1-wt, NPM1-wild-type. Statistically significant difference (^aP<0.05).

Discussion

Previous studies have investigated the prognostic significance of NPM1 mutation status in AML patients. Certain studies demonstrated the positive prognostic effect of NPM1 mutations (13-16,19-22), whereas in other studies no clinical outcome difference between patients with and without NPM1 mutations (17,18). The aim of the present meta-analysis was to clarify the prognostic significance of NPM1 mutation status in AML patients. Meta-analysis is a useful statistical method for integrating results from independent studies for a specified outcome. Combining the relevant studies increases statistical power, thus the effects that may be missed by individual studies may be detected (42). The present meta-analysis demonstrated the effects of NPM1 mutations with the summary OR of 2.23 (95% CI: 1.86-2.67) for CR, HR of 0.67 (95% CI: 0.58-0.78) for DFS, and HR of 0.63 (95% CI: 0.54-0.74) for OS. Furthermore, the present study indicated that NPM1 mutations were associated with a higher frequency of normal karyotype and FLT3-ITD mutations.

Notably, FLT3-ITD mutations have been shown to be the most important abnormality in AML patients and correlated with marked poor outcome (high percentage of bone marrow blast cells, increased risk of relapse from CR, and reduced survival) (43,44). However, our study indicated that patients with FLT3-ITD and NPM1 mutations have an improved CR (14,15,19,20), DFS (14,20) and OS (14,20) compared with those who only have the FLT3-ITD mutation, although this result was inferior to only NPM1 mutation cases.

The present study has several limitations. Firstly, the analyses were based on observational studies rather than prospective controlled studies or randomized trials. Secondly, we used abstracted data, while an individual patient data-based meta-analysis would have provided a more robust estimate of the association. The results reported should therefore be interpreted carefully by clinical physicians. Thirdly, as is often the case with meta-analysis, a substantial effect of heterogeneity should be considered. Although median WBC counts at the time of diagnosis were not identified as sources of heterogeneity, we cannot rule out the potential effect of other factors, such as differences in treatment and distinct cytogenetic categories, which were not examined in our analysis. Publication bias, although not directly detected in the present study, may also have had an impact on the accuracy of our study (42,45).

Although these limitations should be considered, the results of our meta-analysis demonstrated that NPM1 mutations have a favorable effect on the outcome for AML. Thus, distinguishing AML with NPM1 mutations from AML without mutations and justifying the risk-adapted therapeutic strategy for AML based on the NPM1 status.

A large number of patients should be prospectively studied in order for definitive conclusions to be reached. In addition to the presence or absence of NPM1 mutations, several factors relevant to NPM1 have been suggested to have a prognostic value, including the expression levels of NPM1 transcripts, mutant/wild-type allelic ratios for NPM1, and certain types of NPM1 exon mutations. These factors should be investigated in order to arrive at a more accurate estimation of the prognosis for AML.

Acknowledgements

The authors would like to thank Dr Di Wu and Dr Jieying Xi for their technological assistance.

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